

## Solution exercise set 3

### Exercises from the book:

#### 6.40

$X$ =consumption of water is having a gamma distribution with  $\alpha = 2$  and  $\beta = 3$ .

$$\begin{aligned} P(X > 9) &= \int_9^{\infty} \frac{1}{3^2\Gamma(2)} x^{2-1} e^{-x/3} dx = \frac{1}{9} \int_9^{\infty} x e^{-x/3} dx \\ &= \frac{1}{9} [x(-3)e^{-x/3}]_9^{\infty} - \frac{1}{9} \int_9^{\infty} -3e^{-x/3} dx \\ &= \frac{1}{9} 27e^{-3} + \frac{1}{9} [-9e^{-x/3}]_9^{\infty} = 3e^{-3} + e^{-3} = 4e^{-3} = \underline{\underline{0.199}} \end{aligned}$$

Here integration by parts is used to solve the integral. Alternatively the integration formula found on the last page of the formula sheets could have been used.

#### 6.43

a) By the general formulas for expectation and variance in the gamma distribution we get:

$$E(X) = \alpha\beta = 2 \cdot 3 = \underline{\underline{6}} \text{ and } \text{Var}(X) = \alpha\beta^2 = 2 \cdot 3^2 = \underline{\underline{18}}.$$

**6.47**

a) For the Weibull distribution with parameters  $\alpha$  and  $\beta$  we have that (see the paragraph about the gamma function on the formula sheets)

$E(X) = \alpha^{-1/\beta} \Gamma(1 + \frac{1}{\beta})$ . I.e.

$$E(X) = 0.5^{-1/2} \Gamma(1 + \frac{1}{2}) = \sqrt{2} \cdot \Gamma(\frac{3}{2}) = \sqrt{2}(\frac{3}{2} - 1) \Gamma(\frac{1}{2}) = \sqrt{\frac{\pi}{2}} = \underline{\underline{1.25}}$$

b)

$$f(x) = \alpha \beta x^{\beta-1} e^{-\alpha x^\beta} = x e^{-\frac{1}{2}x^2}, x \geq 0 \quad \text{which gives that:}$$

$$P(X > 2) = \int_2^\infty x e^{-\frac{1}{2}x^2} dx = [-e^{-\frac{1}{2}x^2}]_2^\infty = e^{-\frac{1}{2}2^2} = e^{-2} = \underline{\underline{0.135}}$$

**6.56/6.54**

We use the fact that when  $X$  is having a lognormal distribution then  $\ln(X)$  is having a normal distribution and get:

$$\begin{aligned} P(X > 270) &= 1 - P(X \leq 270) = 1 - P(\ln(X) \leq \ln(270)) \\ &= 1 - P\left(\frac{\ln(X) - 4}{2} \leq \frac{\ln(270) - 4}{2}\right) \\ &= 1 - P(Z \leq 0.80) = 1 - 0.7881 = \underline{\underline{0.2119}} \end{aligned}$$

**6.57/6.55**

For the lognormal distribution we have that  $E(X) = e^{\mu+\sigma^2/2}$  and  $\text{Var}(X) = e^{2(\mu+\sigma^2)} - e^{2\mu+2\sigma^2}$  which here gives that:

$$\begin{aligned} E(X) &= e^{4+2^2/2} = e^6 = \underline{\underline{403.4}} \\ \text{Var}(X) &= e^{2(4+2^2)} - e^{2 \cdot 4 + 2^2} = e^{16} - e^{12} = \underline{\underline{8723355.7}} \end{aligned}$$

**6.58/6.56**

a) Let  $X$  be the number of cars arriving at the intersection during one minute. From the information in the exercise text we get that  $X$  is having a Poisson distribution with expectation 5 such that:

$$P(X > 10) = 1 - P(X \leq 10) \stackrel{\text{table}}{=} 1 - 0.9863 = \underline{\underline{0.0137}}$$

b) Here we shall calculate the probability that it will take more than 2 minutes before 10 cars have appeared at the intersection. This can be calculated in two ways.

One (the simplest) way to calculate this is to define  $Y$  as the number of events in the interval  $[0, 2]$  and calculate  $P(Y < 10)$  (the event that it takes more than 2 minutes until 10 cars have appeared is the same as the event that less than 10 cars appear during the 2 minutes).  $Y$  is having a Poisson distribution with expectation  $\lambda t = 5 \cdot 2 = 10$  and we get:

$$P(Y < 10) = P(Y \leq 9) \stackrel{\text{table}}{=} \underline{\underline{0.458}}$$

The other way is to define  $S_{10}$  as the time which elapses until car number 10 appears at the intersection and calculate  $P(S_{10} > 2)$ . We know that time until event number 10 in a Poisson process with  $\lambda = 5$  is having a gamma distribution with parameters  $\alpha = 10$  and  $\beta = 1/5$  (being the sum of 10 exponentially distributed variables with expectation  $1/5$ ). We then get that:

$$\begin{aligned} P(S_{10} > 2) &= 1 - P(S_{10} \leq 2) = 1 - \int_0^2 \frac{1}{\beta^\alpha \Gamma(\alpha)} s^{\alpha-1} e^{-s/\beta} ds \\ &= 1 - \int_0^2 \frac{1}{(\frac{1}{5})^{10} \Gamma(10)} s^{10-1} e^{-5s} ds \stackrel{u=5s}{=} 1 - \int_0^{10} \frac{5^{10}}{\Gamma(10)} (\frac{u}{5})^{10-1} e^{-u} du / 5 \\ &= 1 - \int_0^{10} \frac{u^{10-1}}{\Gamma(10)} e^{-u} du \stackrel{\text{table A.24}}{=} 1 - 0.542 = \underline{\underline{0.458}} \end{aligned}$$

### 6.59/6.57

The time between events in a Poisson process is having an exponential distribution with expectation  $1/\lambda = 1/5$ .

a)

$$P(X > 1) = 1 - \int_0^1 5e^{-5x} dx = 1 - [-e^{-5x}]_0^1 = 1 + e^{-5} - 1 = \underline{\underline{0.0067}}$$

b)  $E(X) = \beta = 1/5 = \underline{\underline{0.2}}$

### Exercises from the note on extreme values:

#### Exercise 1:

We have a parallel system made up of two independent components, where the lifetime of each component is having an exponential distribution with parameter  $\lambda$ .

$$F_{X_i}(t) = 1 - \exp(-\lambda t)$$

Let the lifetime of the system be denoted  $V$ . We shall find the distribution of  $V$ . The system is functioning as long as at least one of the components are functioning.

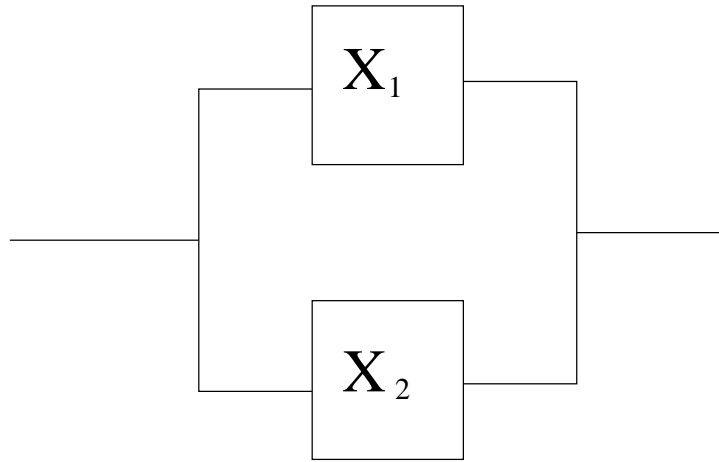


Figure 1: Parallel system with two components.

$$\begin{aligned}
 F_V(v) &= P(V < v) = P(\max(X_1, X_2) < v) = P(X_1 < v \cap X_2 < v) \\
 &\stackrel{\text{indep.}}{=} P(X_1 < v) \cdot P(X_2 < v) = (1 - e^{-\lambda v})^2 \\
 &= 1 - 2e^{-\lambda v} + e^{-2\lambda v} \\
 f_V(v) &= F'_V(v) = 2\lambda e^{-\lambda v} - 2\lambda e^{-2\lambda v} = \underline{\underline{2\lambda(e^{-\lambda v} - e^{-2\lambda v})}}, \quad v \geq 0
 \end{aligned}$$

Expectation:

$$\begin{aligned}
 E(V) &= \int_0^\infty v f_V(v) dv = \int_0^\infty v 2\lambda (e^{-\lambda v} - e^{-2\lambda v}) dv \\
 &= 2 \int_0^\infty v \lambda e^{-\lambda v} - \int_0^\infty v 2\lambda e^{-2\lambda v} = 2 \frac{1}{\lambda} - \frac{1}{2\lambda} = \underline{\underline{\frac{3}{2\lambda}}}
 \end{aligned}$$

Notice that we recognize the two last integrals as the expression for the expectation for exponentially distributed variables with respectively parameter  $\lambda$  and parameter  $2\lambda$ , and we thus know what these integrals are.

### Exercise 2:

We have a series system made up of  $n$  independent components where the lifetime of component  $i$  is having a Weibull distribution with parameters  $\alpha$  and  $\beta$ :

$$P(X_i \leq x) = F_{X_i}(x) = 1 - e^{-\alpha x^\beta}, \quad x \geq 0$$

Let the lifetime of the system be denoted  $U$ . We shall first find the distribution of  $U$ . The system is only functioning as long as all components are functioning.

$$\begin{aligned}
 F_U(u) &= P(U < u) = P(\min(X_1, X_2, \dots, X_n) < u) = 1 - P(\min(X_1, X_2, \dots, X_n) > u) \\
 &\stackrel{\text{indep.}}{=} 1 - \prod_{i=1}^n P(X_i > u) = 1 - \prod_{i=1}^n e^{-\alpha u^\beta} = \underline{\underline{1 - e^{-n\alpha u^\beta}}}
 \end{aligned}$$

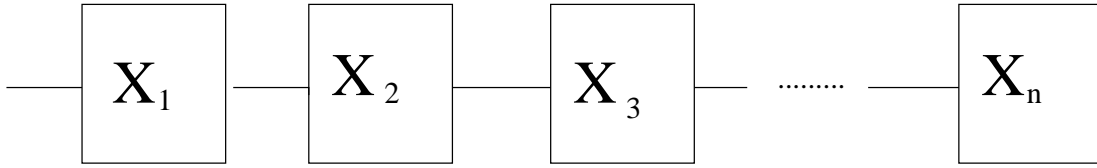


Figure 2: Series system with  $n$  components.

We see that this is a Weibull distribution with parameters  $n\alpha$  and  $\beta$ .

For the Weibull distribution with parameters  $\alpha$  and  $\beta$  we have that  $E(X) = \alpha^{-1/\beta}\Gamma(1 + \frac{1}{\beta})$ .  
 I.e. for a single component we have that:

$$E(X) = 0.1^{-1/0.5}\Gamma(1 + \frac{1}{0.5}) = 100 \cdot \Gamma(3) = 100 \cdot 2! = \underline{\underline{200}}$$

For the whole system we get:

$$E(U) = (n \cdot 0.1)^{-1/0.5}\Gamma(1 + \frac{1}{0.5}) = \frac{1}{n^2}100 \cdot \Gamma(3) = \frac{200}{\underline{\underline{n^2}}}$$

**Exercise 1:**

a)

$$\begin{aligned} E(X) &= \int_{-\infty}^{\infty} xf(x)dx = \int_0^{\infty} x \frac{1}{\beta} e^{-x/\beta} dx = [-xe^{-x/\beta}]_0^{\infty} - \int_0^{\infty} -e^{-x/\beta} dx = 0 - [\beta e^{-x/\beta}]_0^{\infty} = \underline{\underline{\beta}} \\ E(X^2) &= \int_0^{\infty} x^2 \frac{1}{\beta} e^{-x/\beta} dx = [-x^2 e^{-x/\beta}]_0^{\infty} - \int_0^{\infty} 2x(-e^{-x/\beta}) dx \\ &= 0 + 2\beta \int_0^{\infty} x \frac{1}{\beta} e^{-x/\beta} dx = 2\beta^2 \\ \Rightarrow \text{Var}(X) &= E(X^2) - E(X)^2 = 2\beta^2 - \beta^2 = \underline{\underline{\beta^2}} \end{aligned}$$

Instead of integration by parts integration formulas found on the last page of the formula sheets could have been used to solve the integrals.

b) With  $\beta = 1000$  we get

$$P(X > 1000) = \int_{1000}^{\infty} \frac{1}{1000} e^{-x/1000} dx = [-e^{-x/1000}]_{1000}^{\infty} = e^{-1} = \underline{\underline{0.368}}$$

**Exercise 2:**

a) The number of telephone calls per hour is having a Poisson distribution with expectation  $\mu = \lambda t = 6 \cdot 1 = 6$ , i.e.

$$P(X > 6) = 1 - P(X \leq 6) \stackrel{\text{table}}{=} 1 - 0.6063 = \underline{\underline{0.3937}}$$

b) The time until the first event in a Poisson process is having an exponential distribution with expectation  $1/\lambda = 1/6$ . Also note that 10 minutes is  $1/6$  hour. E.g.

$$P(T < 1/6) = \int_0^{1/6} 6e^{-6x} dx = [-e^{-6x}]_0^{1/6} = -e^{-1} + 1 = \underline{\underline{0.6321}}$$

(Alternatively the exercise can be solved by looking at the probability of having at least one event in a 10 minute interval)

c) The time until the second event in a Poisson process is having a gamma distribution with parameters  $\alpha = 2$  and  $\beta = 1/\lambda = 1/6$ . Since 20 minutes is  $1/3$  hour we then get:

$$\begin{aligned} P(S_2 < 1/3) &= \int_0^{1/3} \frac{1}{(1/6)^2 \Gamma(2)} x^{2-1} e^{-x/(1/6)} dx = \int_0^{1/3} 36x e^{-6x} dx \\ &= [-6x e^{-6x}]_0^{1/3} - \int_0^{1/3} (-6e^{-6x}) dx \\ &= -2e^{-2} - [e^{-6x}]_0^{1/3} = -2e^{-2} - e^{-2} + 1 = \underline{\underline{0.5940}} \end{aligned}$$

Alternatively we can solve the problem by defining  $Y$  = the number of telephone calls in  $[0, 1/3]$ ,  $Y$  is then having a Poisson distribution with expectation  $\mu = \lambda t = 6 \cdot (1/3) = 2$ . If it takes less than 20 minutes until the second telephone call, this means that there will be at least 2 calls during the 20 first minutes ( $1/3$  hour) and we get:

$$P(S_2 < 1/3) = P(Y \geq 2) = 1 - P(Y \leq 1) \stackrel{\text{table}}{=} 1 - 0.4060 = \underline{\underline{0.5940}}$$

d) The number of telephone calls during 7.5 hours is having a Poisson distribution with expectation  $\mu = \lambda t = 6 \cdot 7.5 = 45$ . In principle this can be used to calculate  $P(X > 50)$  exactly, however, in this case it is easier to use the approximation to the normal distribution. Since  $\mu > 15$  the approximation to the normal distribution is good, and we get

$$\begin{aligned} P(X > 50) &= 1 - P(X \leq 50) = 1 - P\left(\frac{X - 45}{\sqrt{45}} \leq \frac{50 + 0.5 - 45}{\sqrt{45}}\right) \\ &= 1 - P(Z \leq 0.82) = 1 - 0.7939 = \underline{\underline{0.2061}} \end{aligned}$$

e) The number of events in any interval of length  $t$  in a Poisson process is having a Poisson distribution with expectation  $\lambda t$ , i.e. the number of telephone calls during 10 minutes = 1/6 hour is having a Poisson distribution with expectation  $\mu = 6 \cdot (1/6) = 1$ , i.e.

$$P(X = 0) = \frac{1^0}{0!} e^{-1} = \underline{\underline{0.3679}}$$

f) Since the number of events in non-overlapping intervals in a Poisson process are independent (independent increments) what has happened in the previous 10 minutes has no influence on what will happen the next 10 minutes - i.e. the probability of no telephone calls the next 10 minutes is the same as the probability of no calls in any 10 minute interval, 0.3679.

### Exercise 3:

a) We first find:

$$F_X(x) = \int_0^x 0.02ue^{-0.01u^2} du = [-e^{-0.01u^2}] = 1 - e^{-0.01x^2}$$

Further we have:

$$\begin{aligned} F_U(u) &= P(U < u) = P(\min(X_1, X_2, X_3) < u) = 1 - P(\min(X_1, X_2, X_3) > u) \\ &\stackrel{\text{indep.}}{=} 1 - P(X_1 > u) \cdot P(X_2 > u) \cdot P(X_3 > u) = 1 - [1 - F_X(u)]^3 \\ &= 1 - [e^{-0.01u^2}]^3 = \underline{\underline{1 - e^{-0.03u^2}}} \end{aligned}$$

We could also calculate the probability density:  $f_U(u) = F'_U(u) = \underline{\underline{0.06ue^{-0.03u^2}}}$ ,  $u > 0$

b)

$$\begin{aligned} P(X < 5) &= F_X(5) = 1 - e^{-0.01 \cdot 5^2} = \underline{\underline{0.221}} \\ P(U < 5) &= F_U(5) = 1 - e^{-0.03 \cdot 5^2} = \underline{\underline{0.528}} \end{aligned}$$

c) If we compare the probability density in a) with the Weibull distribution,

$$f(x) = \alpha\beta x^{\beta-1} e^{-\alpha x^\beta}, \quad x \geq 0,$$

we see that  $X$  is having a Weibull distribution with  $\alpha = 0.01$  and  $\beta = 2$ , while  $U$  is having a Weibull distribution with  $\alpha = 0.03$  and  $\beta = 2$ . From the expression for the expectation in the Weibull distribution,  $E(X) = \alpha^{-1/\beta} \Gamma(1 + \frac{1}{\beta})$  we then get

$$\begin{aligned} E(X) &= 0.01^{-1/2} \Gamma(1 + \frac{1}{2}) = 0.01^{-1/2} \frac{1}{2} \Gamma(\frac{1}{2}) = 0.01^{-1/2} \frac{1}{2} \sqrt{\pi} = \underline{\underline{8.86}} \\ E(U) &= 0.03^{-1/2} \Gamma(1 + \frac{1}{2}) = 0.03^{-1/2} \frac{1}{2} \Gamma(\frac{1}{2}) = 0.03^{-1/2} \frac{1}{2} \sqrt{\pi} = \underline{\underline{5.12}} \end{aligned}$$